

Student's t-distribution

Let X and Y be independent random variables where $X \sim N(0,1)$ and $Y \sim \chi_n^2$ then $T = \frac{X}{\sqrt{Y/n}}$ is said to have t-df n on n d.f.

$$T \sim t_n$$

$$f_T(t) = \frac{\Gamma(\frac{n+1}{2})}{\Gamma(\frac{n}{2})\sqrt{\pi n}} (1+t^2/n)^{-\frac{(n+1)}{2}}, -\infty < t < \infty$$

$$= \frac{1}{\sqrt{n} B(\frac{n}{2}, \frac{1}{2})} (1+t^2/n)^{-\frac{(n+1)}{2}}, -\infty < t < \infty$$

The pdf of T is symmetric about 0.
Hence moments of odd order = 0, provided they exist.

For even ordered moment

$$E(T^k) = \frac{n^{k/2} \frac{\Gamma(\frac{k+1}{2}) \Gamma(\frac{n-k}{2})}{\Gamma(\frac{n+1}{2})}}{\Gamma(\frac{n}{2}) \sqrt{\pi n}}, k < n$$

$$E(T) = 0, E(T^2) = V(T) = \frac{n}{n-2}, n > 2$$

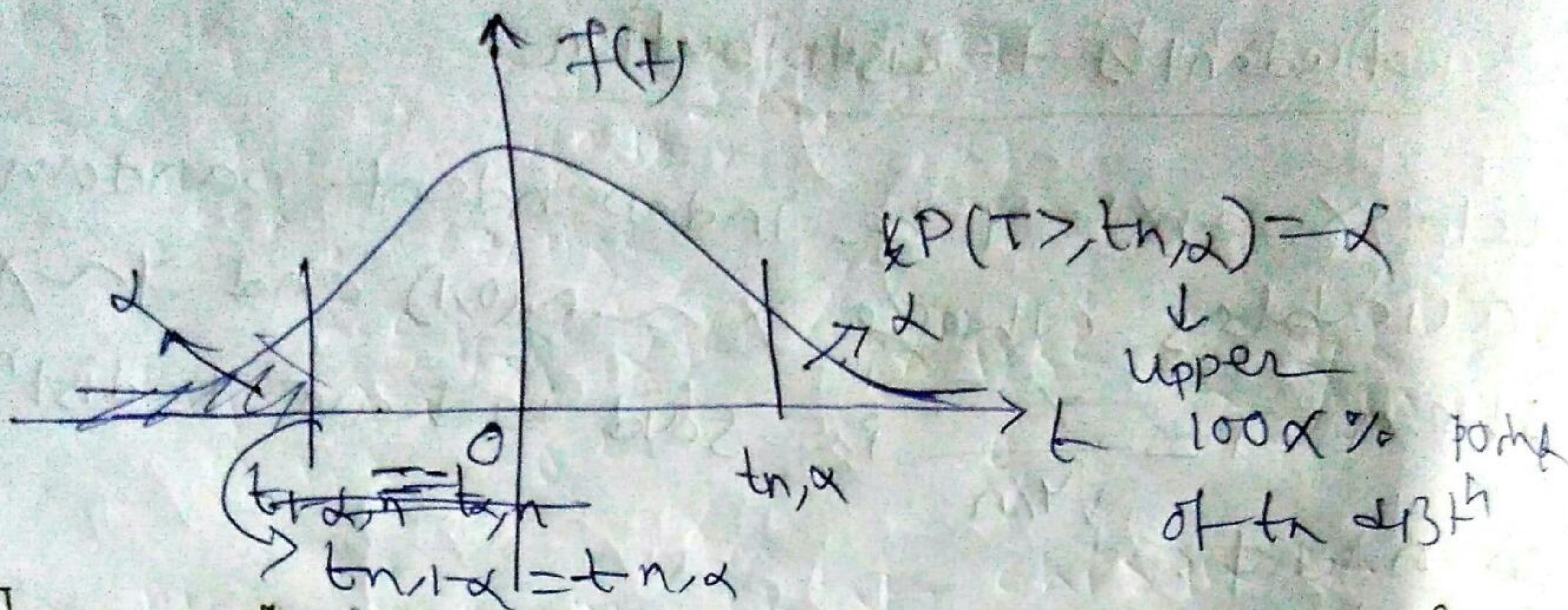
$$M_4 = E(T^4) = \frac{3n^2}{(n-2)(n-4)}, n > 4$$

$$\beta_2 = \frac{M_4}{M_2^2} - 3 = \frac{6}{n-4} > 0$$

↳ measure of kurtosis

So, the density is leptokurtic.

(As $n \rightarrow \infty$ pdf \rightarrow pdf of $N(0,1)$)



Theorem: Let $T \sim t_n$, as $n \rightarrow \infty$ the pdf of T converges to $\phi(t)$.

Remark: For $n > 30$ the approx is very good.

Let $x_1, x_2, \dots, x_n \stackrel{\text{i.i.d}}{\sim} N(\mu, \sigma^2)$.

$$\bar{X} = \frac{1}{n} \sum_{i=1}^n x_i \sim N\left(\mu, \frac{\sigma^2}{n}\right).$$

$$\frac{(n-1)S^2}{\sigma^2} \sim \chi_{n-1}^2 \text{ and } \bar{X} \text{ and } S^2 \text{ are independent}$$

$$\text{So, } \frac{\frac{\sqrt{n}(\bar{X} - \mu)}{\sigma}}{\sqrt{\frac{(n-1)S^2}{\sigma^2(n-1)}}} \sim t_{n-1} \quad \left| \quad \begin{array}{l} \frac{\bar{X} - \mu}{\sigma/\sqrt{n}} \sim N(0,1) \\ \frac{\bar{X} - \mu}{S/\sqrt{n}} \sim t_{n-1} \end{array} \right.$$

$$\therefore \frac{\sqrt{n}(\bar{X} - \mu)}{S} \sim t_{n-1}.$$

Sneddon's f-distribution

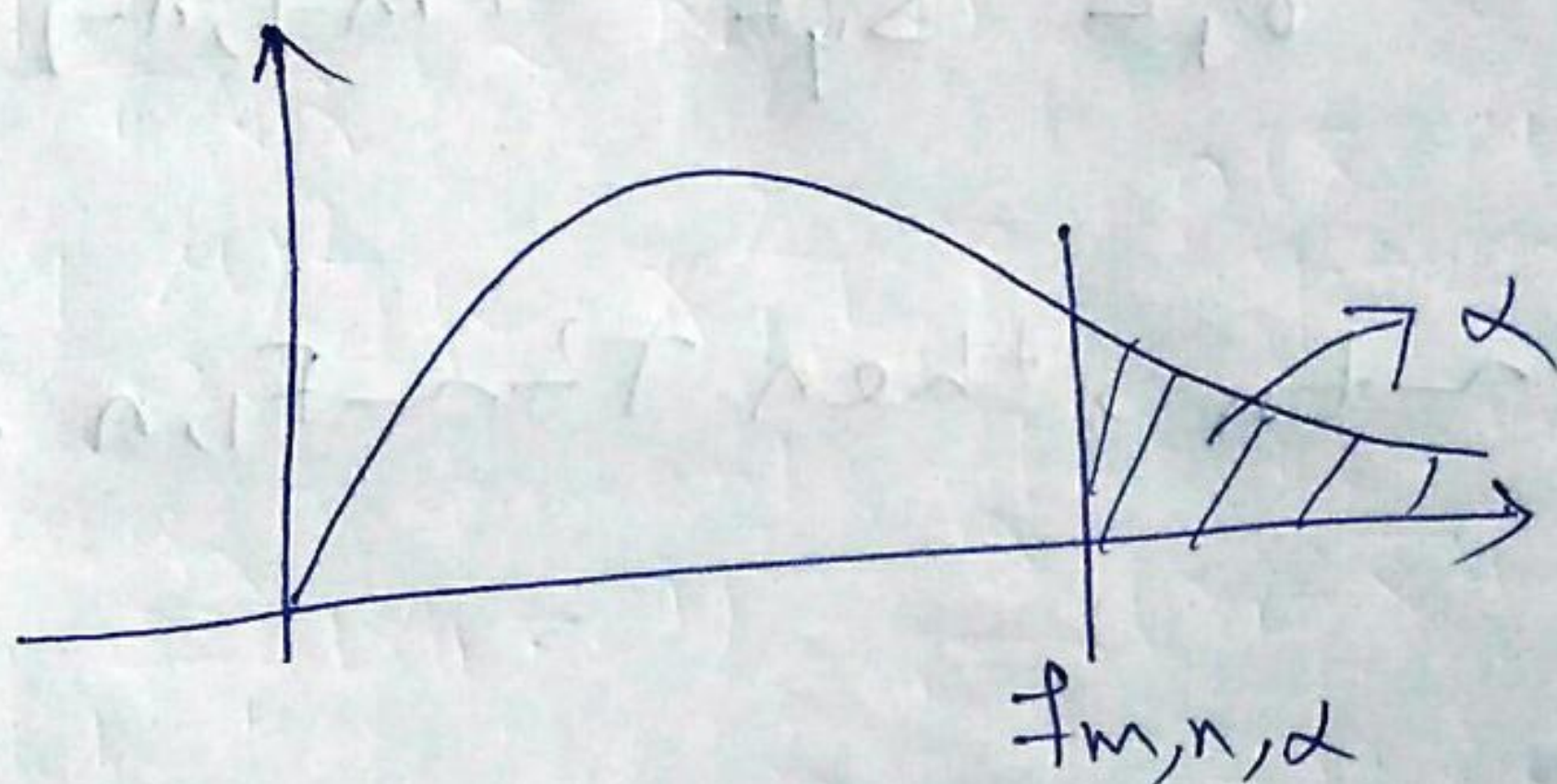
Let Y_1 and Y_2 be independent χ_m^2 and χ_n^2 r.v. Then $W = \frac{Y_1}{Y_2}$

$W = \frac{Y_1/m}{Y_2/n} = \frac{nY_1}{mY_2}$ follows an $F_{m,n}$ dist

$$f_W(w) = \frac{\left(\frac{m}{n}\right)^{m/2} \frac{1}{B\left(\frac{m}{2}, \frac{n}{2}\right)} \frac{w^{m/2-1}}{\left(1 + \frac{m}{n}w\right)^{\frac{m+n}{2}}}; w > 0$$

$$E(W) = \frac{n}{n-2}, n > 2$$

$$V(W) = \frac{2n^2(m+n-2)}{m(n-2)^2(n-4)}, n > 4$$



$$P(W > f_{m,n,\alpha}) = \alpha.$$

• If $U \sim F_{m,n}$ then $\frac{1}{U} \sim F_{n,m}$.

$$\begin{aligned} \text{so } \alpha &= P(U > f_{\alpha,m,n}) = P\left(\frac{1}{U} \leq \frac{1}{f_{\alpha,m,n}}\right) \\ &= P\left(V \leq \frac{1}{f_{\alpha,m,n}}\right) \quad V \sim F_{n,m} \\ &= 1 - P\left(V > \frac{1}{f_{\alpha,m,n}}\right) \end{aligned}$$

$$\text{so, } P\left(V > \frac{1}{f_{\alpha,m,n}}\right) = 1 - \alpha$$

$$\Rightarrow \boxed{f_{1-\alpha,n,m} = \frac{1}{f_{\alpha,m,n}}}$$

Let $X_1, \dots, X_m \stackrel{\text{iid}}{\sim} N(\mu_1, \sigma_1^2)$
~~independent~~ $Y_1, \dots, Y_n \stackrel{\text{iid}}{\sim} N(\mu_2, \sigma_2^2)$

$$S_x^2 = \frac{1}{m-1} \sum_{i=1}^m (x_i - \bar{x})^2 \quad S_y^2 = \frac{1}{n-1} \sum_{j=1}^n (y_j - \bar{y})^2$$

$$U = \frac{(m-1) S_x^2}{\sigma_1^2} \sim \chi_{m-1}^2, \quad V = \frac{(n-1) S_y^2}{\sigma_2^2} \sim \chi_{n-1}^2$$

$$\frac{U/(m-1)}{V/(n-1)} = \frac{\cancel{U}}{\cancel{V}}$$

$$\frac{U/(m-1)}{V/(n-1)} = \frac{\sigma_2^2}{\sigma_1^2} \frac{S_x^2}{S_y^2} \sim F_{m-1, n-1}$$

Thm: If $T \sim t_n$, then $T^2 \sim F_{1, n}$.